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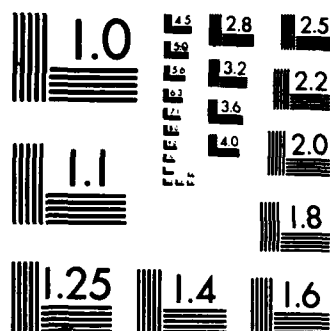
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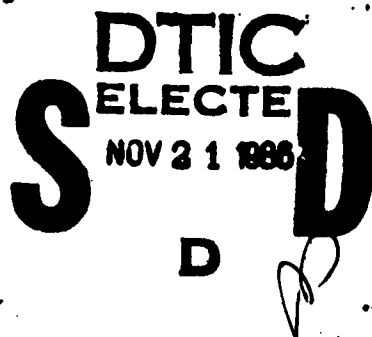




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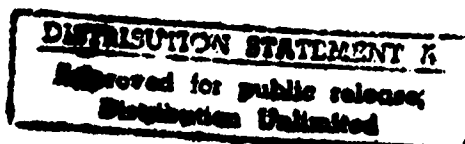
Sheldon M. Ross¹

ORC 86-18

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PEAKS FROM RANDOM DATA

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1. INTRODUCTION

In an influential and controversial paper, Raup and Sepkoski [2]⁹ defined an event of mass extinction to have occurred in any time period (of roughly 6.4 million years) for which the data value for that time period (equal to the proportion of the families existing at the beginning of that period that went extinct during the period) exceeded that of its immediate neighbors. In this note we analyze the occurrence of such events when the data are randomly generated from a continuous distribution.

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2. THE PROCESS OF EVENTS UNDER RANDOM DATA

Let X_1, X_2, \dots be a sequence of independent and identically distributed continuous random variables, and define an event to occur at time period j if

$X_{j-1} < X_j > X_{j+1}$, $j \geq 2$. Let

$$I_j = \begin{cases} 1 & \text{if an event occurs at time } j \\ 0 & \text{otherwise} \end{cases}$$

and set

$$N(n) = \sum_{j=2}^n I_j .$$

In words, $N(n)$ is the number of events by time n .

Proposition 1: With probability 1,

$$\lim_{n \rightarrow \infty} N(n)/n = 1/3$$

Proof: This follows from the ergodic theorem since $\{I_j, j \geq 2\}$ is a stationary ergodic sequence and $E[I_j] = 1/3$. A simpler proof is obtained by noting that $\{I_2, I_5, I_8, \dots\}$, $\{I_3, I_6, I_9, \dots\}$, and $\{I_4, I_7, I_{10}, \dots\}$ are each iid sequences whose average value converges, by the strong law of large numbers, to $1/3$. QED

The mean and variance of $N(n)$ are easily obtained:

$$E[N(n)] = \sum_{j=2}^n E[I_j] = (n-1)/3$$

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$$\text{Var}[N(n)] = \sum_{j=2}^n \text{Var}(I_j) + 2 \sum_{i < j} \text{Cov}(I_i, I_j)$$

As

$$\text{Var}(I_j) = 2/9$$

$$\text{Cov}(I_1, I_{1+1}) = -E[I_1] E[I_{1+1}] = -1/9$$

$$\begin{aligned} \text{Cov}(I_1, I_{1+2}) &= P\{X_{1-1} < X_1 > X_{1+1} < X_{1+2} > X_{1+3}\} - 1/9 \\ &= \int \int \int \int \int_{\substack{x_1 < x_2 > x_3 < x_4 > x_5 \\ 0 < x_i < 1}} dx_1 dx_2 dx_3 dx_4 dx_5 - 1/9 = 1/45 \end{aligned}$$

$$\text{Cov}(I_1, I_{1+j}) = 0, \quad j \geq 3$$

we see from the above that, for $n \geq 3$,

$$\begin{aligned} \text{Var}[N(n)] &= 2(n-1)/9 - 2(n-2)/9 + 2(n-3)/45 \\ &= (2n + 4)/45 \end{aligned}$$

It follows from the above and the central limit theorem for stationary ergodic processes that

$$\frac{N(n) - (n-1)/3}{\sqrt{(2n+4)/45}} \longrightarrow N(0,1)$$

Terminology If $X_{j-1} < X_j > X_{j+1}$ call X_{j+1} the end-of-peak value.

3. INTEREVENT TIMES

Let T_i , $i \geq 1$, denote the time between the i^{th} and $(i+1)^{\text{st}}$ event - for instance, if $I_2=I_4=1$ then $T_1=2$. Even though the T_i are neither independent nor identically distributed we will show that the proportion of them that equal j will, with probability 1, converge to a constant value which we shall call p_j . To determine the values p_j , call an event a j -event if the event preceding it occurred exactly j time periods earlier, and note that by the same argument as used in Proposition 1, the rate at which j -events occur, call it $r(j)$, will equal the limiting probability of a j -event at time period n . Let

$$\bar{R}(k) = \sum_{j \geq k} r(j) .$$

That is, $\bar{R}(k)$ denotes the rate at which interevent times of length greater than k occur. Now, an interevent time will exceed k if, for some $r=0,1,\dots,k$, the next r values after an end of peak are each less, and the next $k-r$ values are each greater, than their preceding values. Therefore, for $k \geq 2$,

$$\begin{aligned} \bar{R}(k) &= P\{X_{-2} < X_{-1} > X_0 < X_1 < \dots < X_k\} + P\{X_{-2} < X_{-1} > X_0 > X_1 > \dots > X_k\} \\ &+ \sum_{r=1}^{k-1} P\{X_{-2} < X_{-1} > X_0 > X_1 > \dots > X_r < X_{r+1} < \dots < X_k\} \\ &= \int_0^1 \frac{(1-x)^k}{k!} \frac{(1-x^2)}{2} dx + \int_0^1 \frac{x^{k+2}}{(k+1)!} dx + \sum_{r=1}^{k-1} \int_0^1 \int_0^x \frac{(x-y)^r}{r!} x \frac{(1-y)^{k-r}}{(k-r)!} dy dx \end{aligned}$$

where the last equality obtains from conditioning the first probability on the value of X_0 , the second on X_{-1} , and the remaining on both X_{-1} and X_r .

As the rate at which events occur is $1/3$ it follows that the proportion of events that are j -events is $3r(j)$. That is,

$$p_j = 3 r(j).$$

A direct integration of the above yields, upon using the identity

$$\sum_{j=2}^k r(j) = 1 - \bar{R}(k) ,$$

that

$$p_2 = 2/5$$

$$p_3 = 1/3$$

$$p_4 = 6/35$$

$$p_5 = 1/15$$

$$p_6 = .02116401$$

and so on. The quantities p_j also represent the limiting distribution of the T_1 . A simple, though tedious, summing of the values $\bar{R}(k)$ yields the (already known) result that the mean of this limiting distribution is 3.

Remark: The above could also have been obtained by consideration of the Markov Chain of successive end-of-peak values. The limiting value of T_1 could then be computed by conditioning on X , the limiting end-of-peak value (whose density is given by $f(x)=3(1-x^2)/2$, $0 < x < 1$), and then computing the probability density of the time between peaks when the initial one has an end-of-peak value X .

References

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2. Raup, D.M. and J. Sepkoski, "Periodicity of Extinctions in the Geologic Past," Proceedings of the National Academy of Sciences, 81, 801-805, 1984.
3. Ross, S.M. "Are Mass Extinctions Really Periodic," Probability in the Engineering and Informational Sciences, January 1987.

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